

Lec 10 : Introduction to Nonparametric Statistics

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Nonparametric model: $y \sim p_\theta$, θ : infinite-dimensional, typically a function
 (typically written as $y \sim p_f$)

Canonical examples:

Regression: given $(x_1, y_1), \dots, (x_n, y_n) \sim p_{XY}$, estimate
 the regression function
 $f(x) := \mathbb{E}[Y | X=x]$

Density estimation: given $x_1, \dots, x_n \sim f$ with an unknown density f ,
 estimate f .

Other examples: in causal inference, interested in:

- causal function: $c(x) = \text{Cov}(Y, W | X=x)$

- conditional/heterogeneous ATE (CATE):

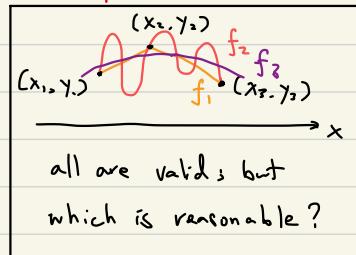
$$\tau(x) = \mathbb{E}[Y | W=1, X=x] - \mathbb{E}[Y | W=0, X=x]$$

Features of nonparametric models:

- model size > sample size; assumptions are necessary to prevent
 overfitting (typically smoothness or shape)

- MLE not well-defined / non-unique / hard
 to find

- explicit bias-variance tradeoff!



Nonparametric regression

A simple binning estimator

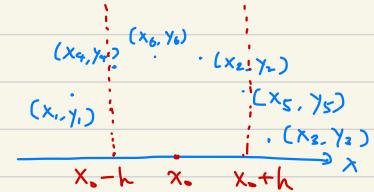
Assumption: $x_1, \dots, x_n \in [0, 1]$ nonrandom (fixed design)

Target: given $x_0 \in [0, 1]$, estimate $f(x_0) = \mathbb{E}[Y | X = x_0]$

bandwidth parameter

A simple estimator:

$$\hat{f}_h(x_0) = \frac{\sum_{i: |x_i - x_0| \leq h} y_i}{\#\{i: |x_i - x_0| \leq h\}}$$



($h \rightarrow 0$: overfit to the closest data point;
 $h \rightarrow \infty$: underfit to the sample average of y)

Analysis: assume that f is L -Lipschitz - i.e. $|f'(x)| \leq L \forall x$

(or equivalently, $|f(x) - f(y)| \leq L|x-y| \forall x, y$)

also, assume that $\text{Var}(Y | X = x) \leq \sigma_0^2$ for all x

Variance of $\hat{f}_h(x_0)$:

$$\begin{aligned} \text{Var}(\hat{f}_h(x_0)) &= \frac{\text{Var}\left(\sum_{i: |x_i - x_0| \leq h} y_i\right)}{\left(\#\{i: |x_i - x_0| \leq h\}\right)^2} \quad (\text{non-random } \{x_i\}) \\ &= \frac{\sum_{i: |x_i - x_0| \leq h} \text{Var}(y_i)}{\left(\#\{i: |x_i - x_0| \leq h\}\right)^2} \quad (\text{independence}) \\ &\leq \frac{\sigma_0^2}{\#\{i: |x_i - x_0| \leq h\}} \end{aligned}$$

If $\{x_i\}$ are evenly spaced in $[0, 1]$, then $\text{Var}(\hat{f}_h(x_0)) = O\left(\frac{\sigma_0^2}{nh}\right)$.

Bias of $\hat{f}_h(x_0)$:

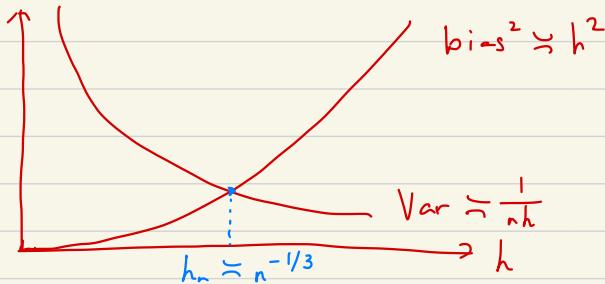
$$\begin{aligned}
 |\text{Bias}(\hat{f}_h(x_0))| &= \left| \mathbb{E}[\hat{f}_h(x_0)] - f(x_0) \right| \\
 &= \left| \frac{\sum_{i: |x_i - x_0| \leq h} f(x_i)}{\#\{i: |x_i - x_0| \leq h\}} - f(x_0) \right| \quad (\mathbb{E}[y_i] = f(x_i)) \\
 &= \left| \frac{\sum_{i: |x_i - x_0| \leq h} (f(x_i) - f(x_0))}{\#\{i: |x_i - x_0| \leq h\}} \right| \\
 &\leq \frac{\sum_{i: |x_i - x_0| \leq h} |f(x_i) - f(x_0)|}{\#\{i: |x_i - x_0| \leq h\}} \quad (\text{triangle inequality}) \\
 &\leq \frac{\sum_{i: |x_i - x_0| \leq h} L|x_i - x_0|}{\#\{i: |x_i - x_0| \leq h\}} \quad (\text{Lipschitz condition}) \\
 &\leq Lh.
 \end{aligned}$$

Final mean-squared error (MSE)

$$\begin{aligned}
 \text{MSE}(\hat{f}_h(x_0)) &= \mathbb{E}(\hat{f}_h(x_0) - f(x_0))^2 \\
 &= \text{Bias}(\hat{f}_h(x_0))^2 + \text{Var}(\hat{f}_h(x_0)) \\
 &= O(L^2 h^2 + \frac{\sigma_0^2}{nh})
 \end{aligned}$$

$$\begin{aligned}
 \text{Optimal choice of } h: h = h_n = \left(\frac{\sigma_0^2}{nL^2} \right)^{1/3} \\
 \text{Optimal MSE} = O(L^{2/3} \sigma_0^{4/3} n^{-2/3})
 \end{aligned}$$

Bias-variance
tradeoff:



Generalization: Nadaraya-Watson estimator

Kernel: a (non-negative) function $K: \mathbb{R}^d \rightarrow \mathbb{R}$ s.t. $\int_{\mathbb{R}^d} K(x) dx = 1$.
Rescaled kernel: for $h > 0$, let $K_h(x) = \frac{1}{h^d} K(\frac{x}{h})$

Examples: rectangle/box kernel: $K(x) = 1(\|x\|_\infty \leq \frac{1}{2})$

Gaussian kernel: $K(x) = (\frac{1}{2\pi})^{d/2} \exp(-\frac{1}{2}\|x\|^2)$

Property: $\int_{\mathbb{R}^d} K_h(x) dx = \int_{\mathbb{R}^d} K_h(hz) h^d dz = \int_{\mathbb{R}^d} K(z) dz = 1, \forall h > 0$.

Kernel-regression estimator / Nadaraya-Watson estimator

$$\hat{f}_h(x_0) = \frac{\sum_{i=1}^n K_h(x_0 - x_i) y_i}{\sum_{i=1}^n K_h(x_0 - x_i)}$$

- previous estimator corresponds to K being the box kernel;
- an equivalent expression of $\hat{f}_h(x_0)$ is

$$\hat{f}_h(x_0) = \sum_{i=1}^n w_i(x_0) y_i, \text{ with } w_i(x_0) = \frac{K_h(x_0 - x_i)}{\sum_{i=1}^n K_h(x_0 - x_i)}$$

being the "weight" of x_i for x_0 .

Analysis. Assume that:

- 1) $|K(x)| \leq B$ for all x ;
- 2) $K(x) = 0$ for all $|x| \geq M$.

$$\begin{aligned} \text{Var}(\hat{f}_h(x_0)) &\leq \frac{\sum_{i=1}^n K_h^2(x_0 - x_i) \sigma_0^2}{\left(\sum_{i=1}^n K_h(x_0 - x_i) \right)^2} \leq \frac{\frac{B}{h^d} \sigma_0^2}{\sum_{i=1}^n K_h(x_0 - x_i)} \\ &= \frac{B \sigma_0^2}{n h^d} \cdot \underbrace{\frac{1}{\frac{1}{nh^d} \sum_{i=1}^n K\left(\frac{x_0 - x_i}{h}\right)}}_{\text{expect to be a constant}} \end{aligned}$$

$$\begin{aligned}
 |\text{Bias}(\hat{f}_h(x_0))| &= \left| \frac{\sum_{i=1}^n K_h(x_0 - x_i) f(x_i)}{\sum_{i=1}^n K_h(x_0 - x_i)} - f(x_0) \right| \\
 &\leq \frac{\sum_{i=1}^n K_h(x_0 - x_i) |f(x_i) - f(x_0)|}{\sum_{i=1}^n K_h(x_0 - x_i)} \quad \begin{array}{l} K_h(x_0 - x_i) = 0 \text{ if} \\ |x_0 - x_i| \geq nh \end{array} \\
 &\leq \frac{\sum_{i=1}^n K_h(x_0 - x_i) \cdot L h M}{\sum_{i=1}^n K_h(x_0 - x_i)} = L M h
 \end{aligned}$$

$$\text{MSE}(\hat{f}_h(x_0)) = O(L^2 M^2 h^2 + \frac{B \sigma^2}{n h^d}) = O(n^{-\frac{2}{2+d}})$$

optimal bandwidth $h = h_n \asymp n^{-\frac{1}{2+d}}$

Capturing higher smoothness of f : next lecture (local polynomials / splines)

Density estimation: estimate f from $x_1, \dots, x_n \sim f$.

Kernels are still useful: let K be a kernel with $\int_{\mathbb{R}^d} K(x) dx = 0$.

Kernel density estimator (KDE):

$$\hat{f}_h(x_0) = \frac{1}{n} \sum_{i=1}^n K_h(x_0 - X_i)$$

- Intuition: when K is the box kernel,

$$\hat{f}_h(x_0) = \frac{\#\{i : x_i \text{ lies in the box centered at } x_0 \text{ of edge length } h\}}{n h^d}$$

$$\approx \frac{f(x_0) n h^d}{n h^d} = f(x_0)$$

Analysis: assume that $\|f''\|_\infty \leq L$, $\int x^2 K(x) dx < \infty$, $\int K^2(x) dx < \infty$, and $d = 1$.

$$\begin{aligned}
 \text{Var}(\hat{f}_h(x_0)) &= \frac{1}{n^2} \sum_{i=1}^n \text{Var}(K_h(x_0 - X_i)) \\
 &\leq \frac{1}{n^2} \sum_{i=1}^n \mathbb{E}[K_h(x_0 - X_i)^2] \\
 &= \frac{1}{nh^2} \int K\left(\frac{x_0 - x}{h}\right)^2 dx \\
 &= \frac{1}{nh} \int K^2(z) dz \quad (z = x_0 - hz)
 \end{aligned}$$

$$\begin{aligned}
 |\text{Bias}(\hat{f}_h(x_0))| &= \left| \mathbb{E}[K_h(x_0 - X_i)] - f(x_0) \right| \\
 &= \left| \underbrace{\int \frac{1}{h} K\left(\frac{x_0 - x}{h}\right) f(x) dx}_{\text{Convolution of } f \text{ and } K_h} - f(x_0) \right| \\
 &= \left| \int \frac{1}{h} K\left(\frac{x_0 - x}{h}\right) (f(x) - f(x_0)) dx \right| \quad (\int K_h(x) dx = 1) \\
 &= \left| \int \frac{1}{h} K\left(\frac{x_0 - x}{h}\right) \left(f(x) - f(x_0) - f'(x_0)(x - x_0) \right) dx \right| \\
 &\quad (\int \frac{1}{h} K\left(\frac{x_0 - x}{h}\right) (x - x_0) dx = 1 \cdot \int z K(z) dz = 0) \\
 &\leq \int \frac{1}{h} K\left(\frac{x_0 - x}{h}\right) \cdot \frac{L}{2} (x_0 - x)^2 dx \\
 &= \frac{Lh^2}{2} \int z^2 K(z) dz \quad (z = x_0 - hz)
 \end{aligned}$$

$$\text{MSE}(\hat{f}_h(x_0)) = O(L^2 h^4 + \frac{1}{nh}) = O(n^{-4/5})$$

↑
optimal bandwidth $h = h_n \approx n^{-1/5}$

View Nadaraya-Watson as KDE:

$$\begin{aligned} \mathbb{E}[Y|X=x] &= \frac{\int y f(x,y) dy}{\int f(x,y) dy} \approx \frac{\int y \cdot \frac{1}{n} \sum_{i=1}^n K_h(x-X_i) K_h(y-Y_i) dy}{\int \frac{1}{n} \sum_{i=1}^n K_h(x-X_i) K_h(y-Y_i) dy} \\ &= \frac{\sum_{i=1}^n K_h(x-X_i) Y_i}{\sum_{i=1}^n K_h(x-X_i)} \end{aligned}$$

Nearest-neighbor density estimator

Define $r_i = \|X_i - x_0\|_2$ as the distance between X_i and x_0 .

For $k=1, 2, \dots, n$, let $r_{(k)}$ be the k -th smallest element of (r_1, \dots, r_n)
(k -th nearest neighbor)

$$\hat{f}_k(x_0) = \frac{k/n}{\text{Vol}_d(r_{(k)})} = \frac{k/n}{\frac{\pi^{d/2}}{\Gamma(\frac{d}{2}+1)} r_{(k)}^d}$$

volume of d -dim ball
of radius $r_{(k)}$

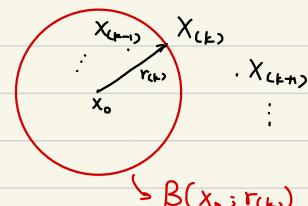
Intuition: $f(x_0) \cdot \text{Vol}_d(r_{(k)})$

$$\approx \int_{B(x_0; r_{(k)})} f(x) dx$$

= actual prob. of $X \in B(x_0; r_{(k)})$

\approx empirical prob. of $X \in B(x_0; r_{(k)})$

$$= \frac{k}{n}.$$



Rigorous claim: $\int_{B(x_0; r_{(k)})} f(x) dx \sim \text{Beta}(k, n-k+1)$

Pf. LHS = k -th smallest element of $\{Z_i = \int_{B(x_0; \|X_i - x_0\|_2)} f(x) dx\}_{i=1}^n$

where $P(Z_i \leq t) = P(\|X_i - x_0\|_2 \leq g^{-1}(t))$ ($g(r) := \int_{B(x_0; r)} f(x) dx$)

$$= g(g^{-1}(t)) = t \Rightarrow Z_i \sim \text{Unif}[0, 1] \quad \square$$

Analysis (Optional) Assume that $\|f'\|_\infty \leq L$ & $c \leq f(x) \leq C$ for all $x \in \text{supp}(f)$.

$$\begin{aligned}
 \text{Step I.} \quad & \left| \int_{B(x_0; r_{(k)})} f(x) dx - f(x_0) \cdot \text{Vol}_d(r_{(k)}) \right| \\
 &= \left| \int_{B(x_0; r_{(k)})} (f(x) - f(x_0)) dx \right| \\
 &= \left| \int_{B(x_0; r_{(k)})} [f(x) - f(x_0) - f'(x_0)(x - x_0)] dx \right| \\
 &\leq \int_{B(x_0; r_{(k)})} \frac{L}{2} \|x - x_0\|_2^2 dx \leq \frac{L}{2} r_{(k)}^2 \text{Vol}_d(r_{(k)})
 \end{aligned}$$

$$\begin{aligned}
 \text{Step II.} \quad & \mathbb{E} \left| \int_{B(x_0; r_{(k)})} f(x) dx - \hat{f}_k(x_0) \cdot \text{Vol}_d(r_{(k)}) \right|^2 \\
 &= \mathbb{E} \left| \text{Beta}(k, n+1-k) - \frac{k}{n} \right|^2 = O\left(\frac{k}{n^2}\right).
 \end{aligned}$$

Step III. Since $f(x) \approx 1$ everywhere,

$$\begin{aligned}
 \frac{k}{n} &\stackrel{\text{w.p.}}{\approx} \int_{B(x_0; r_{(k)})} f(x) dx \approx \text{Vol}_d(r_{(k)}) \\
 \Rightarrow \text{Vol}_d(r_{(k)}) &\approx \frac{k}{n} \Rightarrow r_{(k)} \approx \left(\frac{k}{n}\right)^{1/d}.
 \end{aligned}$$

$$\begin{aligned}
 \text{Conclusion.} \quad f(x_0) \text{Vol}_d(r_{(k)}) &\stackrel{O(r_{(k)} \text{Vol}_d(r_{(k)}))}{\approx} \int_{B(x_0; r_{(k)})} f(x) dx \\
 &\stackrel{O(\sqrt{k/n})}{\approx} \hat{f}_k(x_0) \cdot \text{Vol}_d(r_{(k)}) \\
 \Rightarrow \text{MSE}(\hat{f}_k(x_0)) &= O\left(r_{(k)}^4 + \frac{k/n^2}{\text{Vol}_d(r_{(k)})^2}\right) \\
 &= O\left(\left(\frac{k}{n}\right)^{4/d} + \frac{1}{k}\right) \stackrel{\text{P}}{=} O(n^{-\frac{4}{4+d}}) \\
 k &= k_n \propto n^{\frac{4}{4+d}}
 \end{aligned}$$

(matching the KDE result for $d=1$)