## DS-GA 3001.001 Applied Statistics: Homework #4

Due on Thursday, November 21, 2024

Please hand in your homework via Gradescope (entry code: DKYKGY) before 11:59 PM.

1. Revisit the example of bivariate Gaussian location model we covered in class:

$$\begin{bmatrix} x_1 \\ y_1 \end{bmatrix}, \cdots, \begin{bmatrix} x_n \\ y_n \end{bmatrix} \sim \mathcal{N} \left( \begin{bmatrix} \theta_0 \\ \eta_0 \end{bmatrix}, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} \right),$$

where  $\rho \in [-1, 1]$  is known.

(a) Recall that the estimating equation based on the score for  $\theta_0$  is

$$\frac{1}{n}\sum_{i=1}^{n} \left[ x_i - \widehat{\theta} - \rho(y_i - \widehat{\eta}) \right] = 0.$$

If  $\widehat{\eta} = \eta_0$  is the true nuisance, from the above equation, determine the probability distribution of  $\widehat{\theta} - \theta_0$  which only depends on  $(n, \rho)$ .

- (b) Repeat (a) if  $\hat{\eta} = \eta_0 + \varepsilon$  with a fixed constant  $\varepsilon$ . Your answer should depend on  $(n, \rho, \varepsilon)$ .
- (c) Now consider the efficient score equation

$$\frac{1}{n}\sum_{i=1}^{n}(x_i-\widehat{\theta})=0.$$

Write out the probability distribution of  $\hat{\theta} - \theta_0$ . How does  $\mathbb{E}[(\hat{\theta} - \theta_0)^2]$  compare with (a) and (b)?

2. In the estimation of ATE, in class we modeled the mean outcomes for each group:

$$\mu_0(X) = \mathbb{E}[Y \mid X, W = 0], \quad \mu_1(X) = \mathbb{E}[Y \mid X, W = 1].$$

Another modeling is to model a single mean outcome  $m(X) = \mathbb{E}[Y \mid X]$  and consider the following estimating function

$$f_{(m,e,\tau)}(W,X,Y) = (Y - m(X) - (W - e(X))\tau)(W - e(X)),$$

where  $e(X) = \mathbb{P}(W = 1 \mid X)$  is the propensity score.

- (a) Find the expression of m(X) in terms of  $(\mu_0(X), \mu_1(X), e(X))$ .
- (b) Assuming that  $\mu_1(x) = \mu_0(x) + \tau$  for all x, show that  $f_{(m,e,\tau)}(W,X,Y)$  is a valid estimating function, i.e.

$$\mathbb{E}[f_{(m,e,\tau)}(W,X,Y)] = 0.$$

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(c) Show that  $f_{(m,e,\tau)}(W,X,Y)$  is Neyman orthogonal with respect to (m,e), i.e.

$$\mathbb{E}[\nabla_m f_{(m,e,\tau)}(W,X,Y)] = 0,$$
  
$$\mathbb{E}[\nabla_e f_{(m,e,\tau)}(W,X,Y)] = 0.$$

(d) Show that  $f_{(m,e,\tau)}(W,X,Y)$  is not doubly robust, by arguing that in general

$$\mathbb{E}[f_{(m,\widehat{e},\tau)}(W,X,Y)] \neq 0.$$

3. Consider the same setting for the AIPW estimator in class, but now we aim to estimate the average treatment effect on the treated (ATTE):  $\tau^{\text{ATTE}} = \mathbb{E}[\mu_1(X) - \mu_0(X) \mid W = 1]$ . Consider the following estimating function:

$$f_{(\mu_0, e, \tau^{\text{ATTE}})}(W, X, Y) = \frac{W(Y - \mu_0(X) - \tau^{\text{ATTE}})}{m} - \frac{e(X)(1 - W)(Y - \mu_0(X))}{m(1 - e(X))},$$

where  $e(x) = \mathbb{P}(W = 1 \mid X = x)$  is the propensity score, and  $m = \mathbb{P}(W = 1)$  is the marginal probability of treatment. For simplicity we assume that m is known.

(a) Let p(x) be the pmf of X = x. Using the Bayes rule, show that

$$\mathbb{P}(X = x \mid W = 1) = \frac{p(x)e(x)}{m}.$$

(b) Use (a) to prove the following identity:

$$\tau^{\text{ATTE}} = \mathbb{E}\left[\frac{e(X)}{m}(\mu_1(X) - \mu_0(X))\right].$$

(c) Show that  $f_{(\mu_0,e,\tau^{\text{ATTE}})}(W,X,Y)$  is a valid estimating function, i.e.

$$\mathbb{E}[f_{(\mu_0,e,\tau^{\text{ATTE}})}(W,X,Y)] = 0.$$

(d) (Bonus 5 points) Show that  $f_{(\mu_0,e,\tau^{\text{ATTE}})}(W,X,Y)$  is doubly robust, i.e. for any  $(\widehat{\mu}_0(x),\widehat{e}(x))$ ,

$$\mathbb{E}[f_{(\widehat{\mu}_0, e, \tau^{\text{ATTE}})}(W, X, Y)] = 0,$$
  
$$\mathbb{E}[f_{(\mu_0, \widehat{e}, \tau^{\text{ATTE}})}(W, X, Y)] = 0.$$

- 4. Coding I: we will implement Stein's semiparametric estimator for the symmetric location model  $y_1, \dots, y_n \sim f(y \theta_0)$ , where in our experiment  $f(y) = e^{-|y|}/2$  is the Laplace density. We will experiment on three estimators of  $\theta_0$ :
  - the sample mean of  $(y_1, \dots, y_n)$ ;
  - the MLE with the knowledge of f you should derive the form of the MLE here and find it to be a very simple statistic of  $(y_1, \dots, y_n)$ ;
  - Stein's semiparametric estimator without the knowledge of f.

Based on inline instructions, fill in the missing codes in https://tinyurl.com/5zjf4bzd. Be sure to submit a pdf with your codes, outputs, and colab link.

5. Coding II: we will compare the IPW and AIPW estimators on a synthetic dataset. Based on inline instructions, fill in the missing codes in https://tinyurl.com/y22fams3. Be sure to submit a pdf with your codes, outputs, and colab link.

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